

There are 6 problems, 10 points each. **MAX 60**

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- **Print Your First and Last Name!**
- **Clearly identify your answers to the questions!** No credit for an answer with no justification.

**Problem 1.** We run a sequence of experiments each of which can result in either success or failure. Two consecutive successes are followed by a success with probability 0.8; in all other cases the next experiment is a success with probability 0.5. In the long run, which proportion of the experiments ends in success?

(Begin with setting up a Markov chain and finding its stationary distribution. After that, argue which proportion of the outcomes are a success by conditioning the current outcome on the two previous outcomes.)

**Solution 1:** Let '1' denote a success and '0' denote a failure. Let  $S_j, j = 1, 2, 3, 4$  be the states of the chain, where  $S_1 = 00, S_2 = 01, S_3 = 10, S_4 = 11$ . The matrix of transition probabilities is

	00	01	10	11
00	0.5	0.5	0	0
01	0	0	0.5	0.5
10	0.5	0.5	0	0
11	0	0	0.2	0.8

For the stationary distribution we obtain the equations

$$\pi_1 = 0.5\pi_1 + 0.5\pi_3, \pi_2 = 0.5\pi_1 + 0.5\pi_3, \pi_3 = 0.5\pi_2 + 0.2\pi_4, \pi_4 = 0.5\pi_2 + 0.8\pi_4$$

whence using  $\sum_j \pi_j = 1$ , we get  $\pi_1 = \pi_2 = \pi_3 = 2/11; \pi_4 = 5/11$ . Next let  $X_i$  be the RV that denotes the  $i$ th state of the chain. Using the total probability formula, we obtain for  $i \geq 2$

$$P(X_i = 1) = \sum_{j=1}^4 P\{X_i = 1 | (X_{i-2}X_{i-1}) = S_j\} \cdot P\{(X_{i-2}X_{i-1}) = S_j\}.$$

In the long run,  $P\{(X_{i-2}X_{i-1}) = S_j\} = \pi_j, j = 1, \dots, 4$  so

$$P(X_i = 1) = \pi_1 p_{12} + \pi_2 p_{24} + \pi_3 p_{32} + \pi_4 p_{44} = \frac{2}{11} \frac{3}{2} + \frac{5}{11} \frac{8}{10} = \frac{7}{11} \approx 0.64.$$

**Solution 2:** Make a chain of the following 3 states:

$$S_0 : X_i = *0; \quad S_1 : X_{i-1} = 0, X_i = 1; \quad S_2 : X_{i-1} = X_i = 1$$

Then the transitions are governed by the following matrix:

	$S_0$	$S_1$	$S_2$
$S_0$	0.5	0.5	0
$S_1$	0.5	0	0.5
$S_2$	0.2	0	0.8

The stationary distribution is  $\pi_0 = 4/11; \pi_1 = 2/11; \pi_2 = 5/11$  (we lump together states  $S_1$  and  $S_3$  from Solution 1, making a new state  $S_0$ ).

The proportion of time the chain is in states 1 or 2 equals  $1 - \pi_0 = \frac{7}{11}$ .

**Problem 2.** Let  $X$  be a discrete RV taking integer values. The CDF of  $X$  is given by

$$F_X(x) = \begin{cases} 0, & x < -2 \\ 1/3, & -2 \leq x < 0 \\ 3/4, & 0 \leq x < 3 \\ 1, & x \geq 3. \end{cases}$$

Find  $EX, E(|X|^2), E(X \cdot |X|), E[(X - 2|X|)^2]$  (Hint: begin with finding the PMF of  $X$ ).

**Solution:** We have

$$p_X(-2) = 1/3; p_X(0) = 3/4 - 1/3 = 5/12; p_X(3) = 1/4$$

$$p_{|X|}(2) = 1/3; p_{|X|}(0) = 5/12; p_{|X|}(3) = 1/4$$

Then

$$EX = -2/3 + 3/4 = 1/12$$

$$E(|X|^2) = 4/3 + 9/4 = 43/12.$$

Let  $Z = X \cdot |X|$ , then  $Z$  takes the values  $-4, 0, 9$ , and the pmf of  $Z$  is

$$\begin{array}{ccc} Z & -4 & 0 & 9 \\ p_Z(z) & \frac{1}{3} & \frac{5}{12} & \frac{1}{4} \end{array}$$

Then

$$EZ = -4/3 + 9/4 = 11/12.$$

Finally  $EX^2 = E|X|^2$ , so

$$E[(X - 2|X|)^2] = EX^2 + 4E|X|^2 - 4E(Z) = \frac{5 \cdot 43}{12} - \frac{44}{12} = \frac{57}{4} = 14\frac{1}{4}.$$

Answer:  $1/12; 43/12; -7/12; 20.25$ .

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**Problem 3.** (a) We select 100 independent random real numbers from the set  $\{x : 0 \leq x \leq m\}$ , where  $m$  is a positive integer. These numbers are then rounded down to 3 decimal points. What is the probability that at least 2 of the obtained results equal 0.981?

(b) A Poisson process registers arrivals with rate  $\lambda$  per hour. In the span of 20 hours, what is the probability that there are three non-overlapping one-hour intervals (each starting on the whole hour; not necessarily consecutive) with exactly two arrivals within each of them?

**Solution:** (a) Let  $p = 0.001/m$ , then the probability that each of the numbers falls in the interval  $[0.981, 0.982)$  is  $p$ , and these events are independent. Therefore, using the binomial distribution, the probability that at least 2 numbers are within this interval equals

$$\text{Answer: } 1 - \text{Binom}(100, p, 0) - \text{Binom}(100, p, 1) = 1 - (1 - p)^{100} - 100p(1 - p)^{99}.$$

Or, we can use the Poisson approximation: setting  $\lambda = np = 1000m$ , write

$$\Pr(\geq 2 \text{ numbers}) \approx 1 - e^{-\lambda}(1 + \lambda).$$

(b) Let  $p$  be the probability that there is exactly one arrival within a one-hour period. Then  $p = P(2, 1) = \frac{\lambda^2}{2}e^{-\lambda}$ . Now the answer is given by a binomial probability  $\text{Binom}(20, p, 3)$ .

$$\text{Answer: } \binom{20}{3}p^3(1 - p)^{17} = 1140p^3(1 - p)^{17}.$$

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**Problem 4.** Let  $X$  and  $Y$  have the joint PDF given by

$$f_{XY}(x, y) = 1 \text{ if } 0 \leq x \leq 1, 0 \leq y \leq 1$$

and 0 otherwise.

Find  $P(X + Y \leq 1/2)$ ,  $P(X - Y \leq 1/2)$ ,  $P(XY \leq 1/4)$ ,  $P(X^2 + Y^2 \leq 1)$ .

**Solution:** By computing the areas, we obtain

$$P(X + Y \leq 1/2) = 1/8;$$

$$P(X - Y \leq 1/2) = 7/8,$$

$$P(XY \leq 1/4) = \frac{1}{4} + \int_{1/4}^1 \frac{dx}{4x} = \frac{1}{4}(1 + \ln 4).$$

$$P(X^2 + Y^2 \leq 1) = \pi/4.$$

If you prefer to compute the integrals, then the following is a correct way to write them:

$$\begin{aligned}
 P(X + Y \leq 1/2) &= \int_0^{1/2} \int_0^{\frac{1}{2}-y} dx dy = \frac{1}{2}(y - y^2) \Big|_{1/2}^0 = \frac{1}{8}. \\
 P(X - Y \leq 1/2) &= \int_0^1 \int_{\max\{0, x-\frac{1}{2}\}}^1 dy dx = \frac{1}{2} + \int_{1/2}^1 \int_{x-\frac{1}{2}}^1 dy dx = \frac{1}{2} + \left[ \frac{1}{2} + x \right]_{1/2}^1 = \frac{7}{8} \\
 P(XY \leq 1/4) &= \int_0^1 \int_0^{\min\{1, \frac{1}{4y}\}} dx dy = \int_0^{1/4} \int_0^1 dx dy + \int_{1/4}^1 \int_0^{\frac{1}{4y}} dx dy = \frac{1}{4} + \int_{1/4}^1 \frac{dy}{4y} dx dy \\
 &= \frac{1}{4} + \frac{1}{4} \ln y \Big|_{1/4}^1 = \frac{1}{4}(1 + \ln 4). \\
 P(X^2 + Y^2 \leq 1) &= \int_0^1 \int_0^{\sqrt{1-y^2}} dx dy = \int_0^1 \sqrt{1-y^2} dy \\
 &= \int_0^{\pi/2} \cos^2 t dt = \int_0^{\pi/2} \left( \frac{1}{2} + \frac{1}{2} \cos(2t) \right) dt = \frac{\pi}{4} + \frac{1}{4} \sin(2t) \Big|_0^{\pi/2} = \frac{\pi}{4}.
 \end{aligned}$$

**Problem 5.** We are given  $n$  independent random samples  $X_1, \dots, X_n$  from an unknown distribution of an RV  $X$  with variance  $\text{Var}(X) = 3$ .

(a) Let  $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ . Find  $E\bar{X}_n, \text{Var}(\bar{X}_n)$ .

(b) We guess that the true value of  $EX$  equals  $\bar{X}_n$ . Suppose that our goal is that this guess differs from the true value of  $EX$  by at most 4. What is the probability that the estimate  $\bar{X}_n$  achieves this precision if  $n = 5$ ?

**Solution:** (this is a rephrased Prob. 4 from HW8) Consider the RV  $\bar{X}_n$ , then  $E\bar{X}_n = EX$ ,  $\text{Var}(\bar{X}_n) = \frac{3}{n}$ . Using the Chebyshev inequality,  $P(|\bar{X}_n - EX| \geq 4) \leq \frac{3}{5 \cdot 4^2} = 0.0375$ . The event in question is complementary, so we find the answer to be  $1 - 0.0375 = 0.9625$ .

**Problem 6.** (a) Given  $n$  balls, we place them randomly, one by one, into Urn 1 (with probability  $p$ ) or Urn 2 (with probability  $1 - p$ ). Let  $X_1$  be the number of balls in Urn 1 at the end of the experiment, and let  $X_2 = n - X_1$  be the number of balls in Urn 2. What is the pmf and the mean of  $X_i, i = 1, 2$ ? Are  $X_1$  and  $X_2$  independent?

(b) Now assume that the number of balls  $n$  is a Poisson RV with parameter  $\lambda$ . Once we have a realization of this RV, i.e., the value of  $n$ , we perform the same experiment as in part (a), obtaining the RVs  $X_1$  and  $X_2$ . These RVs are different from those in part (a) because they depend on the realization of  $n$ . Find the PMFs and the means of  $X_1$  and  $X_2$  by conditioning on  $n$  and using the total probability theorem.

**Solution:** (a)  $X_1 \sim \text{Binom}(n, p)$ ,  $EX_1 = np$ ;  $X_2 \sim \text{Binom}(n, 1 - p)$ ,  $EX_2 = n(1 - p)$ . Given the value of  $X_1$  we know exactly the value of  $X_2$ , i.e.,  $P(X_2 = j) \neq P(X_2 = j | X_1 = k)$ , so  $X_1$  and  $X_2$  are dependent.

(b) Using the total probability formula,

$$\begin{aligned}
 P(X_1 = i) &= \sum_{n=0}^{\infty} P(X_1 = i | n) P(n) = \sum_{n=i}^{\infty} \binom{n}{i} p^i (1-p)^{n-i} e^{-\lambda} \frac{\lambda^n}{n!} \\
 &= \left( \frac{p}{1-p} \right)^i e^{-\lambda} \sum_{n=i}^{\infty} \frac{n!}{i!(n-i)!} \frac{((1-p)\lambda)^n}{n!} \\
 &= \left( \frac{p}{1-p} \right)^i e^{-\lambda} \frac{((1-p)\lambda)^i}{i!} \sum_{n=i}^{\infty} \frac{((1-p)\lambda)^{n-i}}{(n-i)!} = \frac{(p\lambda)^i}{i!} e^{-\lambda} e^{(1-p)\lambda} \\
 &= \frac{(p\lambda)^i}{i!} e^{-p\lambda}, \quad i = 0, 1, \dots
 \end{aligned}$$

**Answer:**  $X_1 \sim \text{Poisson}(p\lambda)$ ,  $EX_1 = p\lambda$  and similarly  $X_2 \sim \text{Poisson}((1-p)\lambda)$ ,  $EX_2 = (1-p)\lambda$ .

Or, you can argue that we are splitting the Poisson process for  $n$  into two streams, obtaining two Poisson processes with rates  $p\lambda$  and  $(1-p)\lambda$ .

(It is an interesting question to figure out whether in this case  $X_1$  and  $X_2$  are independent, by computing  $P(X_1 = i | X_2 = k)$ .)